Supplementary Online Contents

Supplementary Methods

Study populations. Statistical analyses, multiple testing correction. Mendelian randomization analysis

Supplementary Tables

Table S1. Genetic variants for the SHBG gene score and their associations with type 2 diabetes.

Table S2. Prospective association magnitudes of SHBG for incident type 2 diabetes in four studies.

Supplementary Figures

Figure S1. Pleiotropy assessment for the individual SNPs and gene scores associated with circulating SHBG.

Figure S2. Cross-sectional associations of SHBG with systemic lipid and metabolite measures for men and women; additional adjustment for insulin and testosterone.

Figure S3. Cross-sectional associations of SHBG with systemic lipid and metabolite measures in individual cohorts.

Figure S4. Associations of individual genetic variants and the SHBG gene scores with potential confounders.

Figure S5. Correspondence between cross-sectional associations and causal effect estimates based on Gene Score A.

Figure S6. Correspondence between cross-sectional associations and causal effect estimates based on Gene Score B.

Supplementary Methods

Study populations

The Northern Finland Birth Cohorts (NFBC1986 and NFBC1966)

The Northern Finland Birth Cohorts 1986 and 1966 were initiated to study factors affecting preterm birth and subsequent morbidity in the two northernmost provinces in Finland. For NFBC1986, the number of deliveries in the birth cohort was 9362, which was 99% of all the deliveries taking place in the target period (July 1985-June 1986). Data collection in 2001-2002 included clinical examination and serum sampling at the age of 15-16 years for 6621 adolescent boys and girls; data from this time point were used for the secondary analyses of the present study.^{1,2} Attendees in the 16-year field study (71%) were representative of the original cohort.² Individuals using oral contraceptives were excluded. In total, 5099 adolescents had the systemic NMR-based metabolite profile measured, of which 95% of the serum samples were drawn after an overnight fast.

The NFBC1966 included 12 058 children, comprising 96% of all births during 1966 in the region.^{3,4} The collection in 1997 included clinical examination and serum sampling at the age of 31 years for 6007 individuals. Data from this time point were analysed in the present study. Attendees in the 31-year field study (52% of target population) were representative of the original cohort.³ Pregnant women and those using oral contraceptives were excluded. In total, 4841 persons had their systemic NMR-based metabolite profile measured, of which 96% of the serum samples were drawn after an overnight fast.⁵

C-reactive protein and plasma insulin were measured by standard clinical assays. SHBG was analysed by time-resolved fluoroimmunoassay (AutoDelfia, PerkinElmer, Turku, Finland). The intra- and inter-assay coefficients of variation for the assay were 3.9 and 4.6 % at 101 nmol/l, respectively. Serum testosterone was measured by an automated chemiluminescence system (ACS-180, Ciba-Corning, Inc., Medfield, MA, USA). Informed written consent was obtained from all participants. The research protocols were approved by the Ethics Committee of Northern Ostrobotnia Hospital District, Finland.

The Cardiovascular Risk in Young Finns Study (YFS)

The Cardiovascular Risk in Young Finns Study was designed to study associations of childhood risk factors to cardiovascular disease in adulthood.⁶ The baseline study conducted in 1980 included 3596 children and adolescents at the ages of 3–18 years. In this study we used data from the 2001 survey that included 2247 individuals with an overnight fasting NMR-based systemic metabolite profile (response rate 63%). These individuals were representative of the baseline cohort.⁶ Pregnant women and those using oral contraceptives were excluded. In total, data for 1870 individuals who had the metabolic profiles measured were used in the study; of those 1377 individuals also had information at 6-year follow-up.

C-reactive protein and plasma insulin were measured by standard clinical assays. SHBG was measured by a Spectria SHBG IRMA kit (Orion Diagnostica, Espoo, Finland) and total testosterone

by a Spectria Testosterone kit (Orion Diagnostica, Espoo, Finland). All participants gave written informed consent, and the local ethics committees of the study sites approved the study.

The National Finnish FINRISK study

The FINRISK 1997 study was conducted to monitor the health of the Finnish population among persons at the ages of 25–74 years at the recruitment.⁷ In total, 8444 individuals were recruited to represent the general population of the study areas. Pregnant women and those using oral contraceptives were excluded from the analyses of the present study. The NMR-based systemic metabolite profiling from serum samples were available for 7200 individuals. The median fasting time was 5h (interquartile range 4–6h). C-reactive protein and plasma insulin were assayed by standard clinical assays. Participants gave written informed consent and the FINRISK study was approved by the Coordinating Ethical Committee of the Helsinki University Hospital District, Helsinki, Finland.

Statistical analyses, multiple testing correction

The rationale for the modified Bonferroni correction is as follows: 1) Bonferroni adjustment is designed for independent tests, i.e., the outcome of one test is not correlated with the others. 2) By breaking a dataset into principal components (PCs), we effectively create new variables that, by definition, are not correlated.⁸ The assumption here is that the independence of the PC scores is equivalent to the degree of freedom of the original dataset — with >95% variance explained, this can only cause a small inaccuracy. This logic also works for the extreme cases: for perfectly correlated variables only one needs to be tested, hence the Bonferroni correction is for one independent test. If all the variables are orthogonal, then PCA will need a large number of components to explain most of the variance, and the Bonferroni correction approaches the full number of dimensions.

Mendelian randomization analysis

Mendelian randomisation analysis using ratio estimator

The causal effect estimates for the metabolic measures were calculated via the ratio estimator, and the corresponding standard errors via the delta method using the following formulas: ^{9,10}

$$\beta_{IV} = \frac{\beta_{GS-metabolite}}{\beta_{GS-SHBG}}$$

$$se_{IV} = abs(\beta_{IV}) \sqrt{\left(\frac{se_{GS-SHBG}}{\beta_{GS-SHBG}}\right)^{2} + \left(\frac{se_{GS-metabolite}}{\beta_{GS-metabolite}}\right)^{2}}$$

where *abs* refers to an absolute value, $\beta_{GS-metabolite}$ is the effect size and $se_{GS-metabolite}$ is the standard

error of the association between the gene score and the metabolic measure meta-analysed using a fixed-effect model across NFBC1966, YFS and FINRISK (n=10 895). $\beta_{GS-SHBG}$ denotes the effect size and $se_{GS-SHBG}$ the corresponding standard error of the association between the gene score and the circulating SHBG level meta-analysed using a fixed-effect model across NFBC1966 and YFS (n=5860).

Mendelian randomisation analysis using summary statistics

The causal effect estimates and corresponding standard error for type 2 diabetes were calculated via the method proposed by Burgess et al.,¹¹ using the following formulas:

$$\beta_{IV} = \frac{\sum_k X_k Y_k \sigma_{Yk}^{-2}}{\sum_k X_k^2 \sigma_{Yk}^{-2}}$$

$$se_{IV} = \sqrt{\frac{1}{\sum_k X_k^2 \sigma_{Yk}^{-2}}}$$

where X_k is the effect size of the association between the genetic variant *k* and SHBG obtained from the SHBG-GWAS study;¹² Y_k is the log-odds ratio and σ_{Yk} the standard error between the genetic variant *k* and type 2 diabetes for 110 452 individuals (cases/controls=26 488/83 964) reported by DIAGRAM consortium.¹³

The causal estimates based on the two multi-SNP gene-scores in the DIAGRAM consortium were further meta-analysed with effect estimates previously published by Ding *et al*¹⁴ (based on rs6257 and rs6259) and by Perry *et al*¹⁵ based on rs1799941 near the *SHBG* gene. Only studies not part of the DIAGRAM consortium data were part included in the meta-analysis of the causal effects. Details and clinical characteristic of the additional cohorts included in the meta-analysis have been described in the original publications.^{14,15}

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						Association with type 2 diabetes ^c		
	SNP	Gene	Chromosome	Effect ^a	Weight ^b	OR (95%CI)	Р	Ν
				allele				
1	rs780093	GCKR	2	С	0.032	1.05 (1.02-1.08)	0.00014	104,493
2	rs17496332	PRMT6	1	G	0.028	0.98 (0.96-1.01)	0.23	104,087
3	rs440837	ZBTB10	8	G	0.028	0.99 (0.96-1.02)	0.45	104,505
4	rs7910927	<i>JMJD1C</i>	10	G	0.048	0.99 (0.97-1.02)	0.65	104,575
5	rs4149056	SLCO1B1	12	Т	0.029	1.00 (0.96-1.04)	1.00	104,229
6	rs8023580	NR2F2	15	С	0.030	1.03 (1.00-1.06)	0.029	99,991
7	rs2411984	ZNF652	17	А	0.033	0.98 (0.95-1.01)	0.28	94,519
8	rs1573036	TDGF3	23 (X)	Т	0.028			
9	rs12150660	SHBG	17	Т	0.082	0.97 (0.93-1.01)	0.16	72,377
10	rs6258	SHBG	17	С	0.261	0.90 (0.78-1.04)	0.14	63,390
11	rs1641537	SHBG	17	С	0.081	0.99 (0.95-1.02)	0.48	96,111
12	rs1625895	SHBG	17	С	0.052	0.99 (0.94-1.03)	0.55	95,557

Table S1. Genetic variants for the SHBG gene score and their associations with type 2 diabetes.

a: The effect allele is the SHBG-raising allele.

b: Weights of the individual genetic variants are based on a recent genome-wide analyses study.¹²

c: Associations between the individual genetic variants and the risk of type 2 diabetes are reported by the DIAGRAM consortium; results for the variant rs1573036 in the X chromosome is not available.¹³

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Publication ^a	Study design	Age at	Follow-	Baseline	Incident	Log ^c	HR (95%CI) ^a
		baseline ^b	up(v)	number	diabetes		
		basenne	up (y)	number	diabetes		
					number		
Women							
Ding et al, 2009 ¹⁴	Nested case-	60.3 (6.1)	10	718	359	Yes	0.40 (0.31-0.51)
	control						
Chen et al, 2012 ¹⁶	Nested case-	50-79	5.9	1,928	642	Yes	0.48 (0.43-0.54)
	control						
					Overall (Women)		0.45 (0.38-0.54)
Men							
Ding et al, 2009 ¹⁴	Nested case-	63.7 (7.6)	8	340	170	Yes	0.43 (0.31-0.59)
	control						
Lakshman et al, 2010 ¹⁷	Cohort	40-70	13	1,128	90	No	0.50 (0.35-0.70)
Vikan et al, 2010 ¹⁸	Cohort	59.4 (10.3)	9.1	1,454	76	No	0.55 (0.39-0.79)
					Overall (Men)		0.49 (0.40-0.59)

Table S2. Prospective association magnitudes of SHBG for incident type 2 diabetes in four studies.

Overall 0.47 (0.41-0.53)

a: Publications were searched from PubMed with key words "SHBG" and "type 2 diabetes" with a search criterion the publication year 2000 or later. The prospective studies were included if the results were reported as risk of type 2 diabetes per 1-SD increment SHBG. All studies listed here are based on independent cohorts.

b: Age at baseline is given as the range of age or mean (SD).

c: Log indicates whether the SHBG concentration was natural log-transformed or not.

d: When there are multiple adjusted models, the hazard ratio is from the model in which least covariates were used.



Figure S1. Pleiotropy assessment for the individual SNPs and gene scores associated with circulating SHBG. The associations between individual SNPs and the 72 metabolic measures are meta-analyses for 10 895 individuals from the NFBC1966, YFS and FINRISK. The associations were adjusted for sex, age and BMI, and calculated as SD-units per each SHBG-raising allele. Associations with P<0.002 are marked with an asterisks. The associations between the gene scores and the metabolic measures are also provided as a reference (SD-units per 1-SD gene score). Gene Score A is composed of 11 SHBG-correlated SNPs after excluding the pleiotropic *GCKR* variant. Gene Score B is composed of 4 SNPs within the *SHBG* gene, including rs12150660, rs6258, rs1641537 and rs1625895.



Figure S2. Cross-sectional associations of SHBG with systemic lipid and metabolite measures for men and women; additionally adjusted for insulin and testosterone. The associations were meta-analysed for 6475 individuals from the NFBC1966 and YFS. The associations were adjusted for insulin and testosterone, in addition to age and body mass index.



Figure S3. Cross-sectional associations of SHBG with systemic lipid and metabolite measures in individual cohorts. The associations between SHBG and the metabolic measures were further replicated in 4834 adolescents from the NFBC1986. Women and men were combined to obtain a clear visualisation. The associations were adjusted for sex, age and body mass index.



Figure S4. The associations of individual genetic variants and the SHBG gene scores with potential confounders. The associations were assessed for 10 895 individuals across three cohorts. $P \ge 0.002$ for all the associations. The large confidence interval estimated for rs6258 is due to its low minor allele frequency (<1%). The genetic variants included in Gene score A and B are as described in the legend of Figure S1.



[SD increment in metabolic measure per 1–SD SHBG]

Figure S6. The correspondence between cross-sectional associations and causal effect estimates based on Gene Score A. The causal estimates were calculated via Mendelian randomization analyses using the gene score as the instrument. Gene Score A is composed of 11 established genetic variants for elevating SHBG (the *GCKR*-variant excluded due to its wide pleiotropic effects). Each point represents a metabolic measure. The horizontal grey lines behind each point denote the confidence intervals for the cross-sectional associations and the vertical grey lines indicate the confidence intervals for the causal effect estimates. A linear fit of the overall correspondence was made to summarize the match between observational (cross-sectional) associations and causal estimates, and R^2 denotes the goodness of the fit. Since the observational and genetic associations were scaled to the same units (SD increment of metabolite concentration per 1-SD SHBG), a slope of ~1 would be expected for a match between observational and causal effect estimates.



[SD increment in metabolic measure per 1–SD SHBG]

Figure S7. The correspondence between cross-sectional associations and causal effect estimates based on Gene Score B. The causal estimates were calculated via Mendelian randomization analyses using the gene score as the instrument. Gene Score B is composed of the 4 genetic variants at the *SHBG*-encoding gene (rs12150660, rs6258, rs1641537 and rs1625895). Each point represents a metabolic measure. The horizontal grey lines behind each point denote the confidence intervals for the cross-sectional associations and the vertical grey lines indicate the confidence intervals for the causal effect estimates. A linear fit of the overall correspondence was made to summarize the match between observational (cross-sectional) associations and causal estimates, and R^2 denotes the goodness of the fit. Since the observational and genetic associations were scaled to the same units (SD increment of metabolite concentration per 1-SD SHBG), a slope of ~1 would be expected for match between observational and causal effect estimates.